

Math 5760/6890: Introduction to Mathematical Finance

Binomial Pricing Models

See Petters and Dong 2016, Section 5.1

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The overall goal:

Given a security's time-0 price $S(0)$, construct a probabilistic model for $S(t)$, $t > 0$.

We do this with discrete time steps:

- Divide $[0, T]$ into $n \in \mathbb{N}$ equally sized intervals, $[t_j, t_{j+1}]$ for $j = 0, \dots, N$.
- $t_j = jh$ with $h = T/n$
- Let $S_j = S(t_j)$
- Model $S_j \mapsto S_{j+1}$ as a multiplicative process:

$$S_{j+1} = G_{j+1}S_j = \begin{cases} uS_j, & \text{with probability } p \\ dS_j, & \text{with probability } 1 - p \end{cases}$$

- We assume (p, u, d) satisfies $p \in (0, 1)$, and $d < 1 < u$.

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- We assume (p, u, d) satisfies $p \in (0, 1)$, and $d < 1 < u$.
- The behavior over the entire period corresponds to an accumulation of multiplicative gross returns G_j :

$$S_n = S_0 \prod_{j=1}^n G_j, \quad L := \log \frac{S_n}{S_0} = \sum_{j=1}^n \log G_j = \sum_{j=1}^n L_j$$

- The random variables $\{G_j\}_{j=1}^n$ are iid (shifted) Bernoulli random variables (Same for $\{L_j\}_{j=1}^n$.)
- Today: we'll describe more details about this model.

A digression: Binomial random variables

D16-S03(a)

Let $X \sim \text{Bernoulli}(p)$, and consider n iid copies of X , $\{X_j\}_{j=1}^n$.

The random variable X satisfies,

$$X = \begin{cases} 1, & \text{with probability } p \\ 0, & \text{with probability } 1 - p \end{cases}$$

X is the outcome of a (binary) coin toss, where the outcome is biased if $p \neq 1/2$.

A digression: Binomial random variables

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$$Y = \sum_{j=1}^n X_j,$$

which counts the number of +1 outcomes (say “heads” outcomes) of the X_j .

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Evidently, Y is a discrete random variable, with outcomes $\{0, 1, \dots, n\}$.

Y is called a **Binomial**(n, p) random variable, and it has a Binomial distribution.

A digression: Binomial random variables

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The random variable X satisfies,

$$X = \begin{cases} 1, & \text{with probability } p \\ 0, & \text{with probability } 1 - p \end{cases} \quad \mathbb{E}X = p \cdot 1 + (1-p) \cdot 0 = p$$

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Some simple first- and second-order statistics are immediately computable:

$$\mathbb{E}Y = \sum_{j=1}^n \mathbb{E}X_j = \sum_{j=1}^n p = np, \quad \text{Var}Y \stackrel{(*)}{=} \sum_{j=1}^n \text{Var}X_j = np(1 - p)$$

where (*) uses the fact that the variance of the sum of independent random variables is the sum of their variances.

$$\begin{aligned}\text{Var } X &= \mathbb{E}(X-p)^2 = p(1-p)^2 + (1-p)(0-p)^2 \\ &= p(1-p) [1] = p(1-p)\end{aligned}$$

The Binomial distribution

D16-S04(a)

The distribution (mass function) of Y is,

$$p_Y(k) = \binom{n}{k} p^k (1-p)^{n-k}, \quad 0 \leq k \leq n$$

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}, \quad 0! = 1$$

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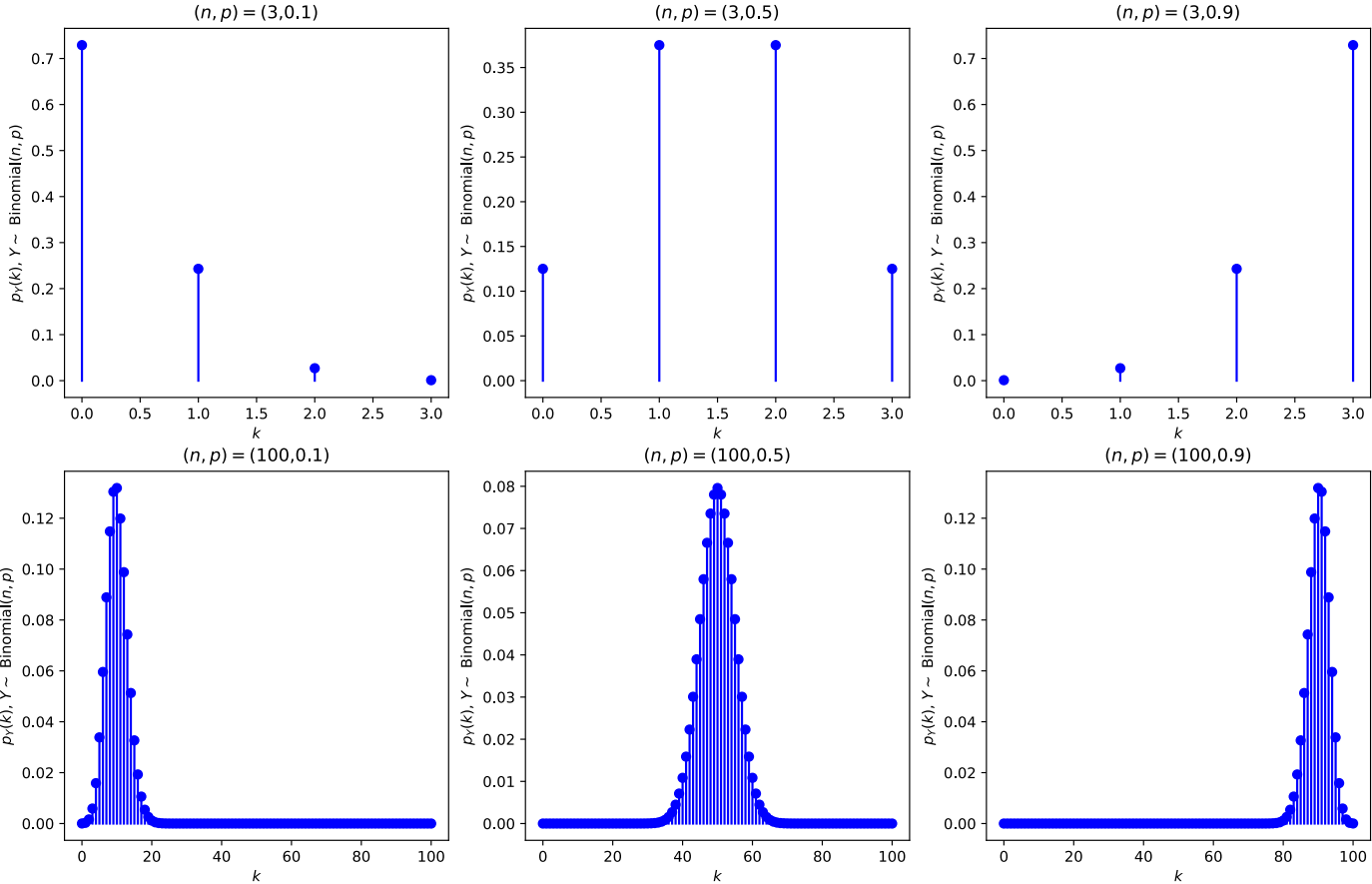
$$p_Y(k) = \binom{n}{k} p^k (1-p)^{n-k}, \quad 0 \leq k \leq n$$

This is a valid mass function due to the Binomial Theorem:

$$\begin{aligned} (x + y)^n &= \sum_{k=0}^n \binom{n}{k} x^k y^{n-k} \quad \implies \quad \sum_{k=0}^n p_Y(k) = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} \\ &= (p + (1-p))^n = 1. \end{aligned}$$

If $Y \sim \text{Binomial}(n, p)$, the quantity $p_Y(k)$ is the probability that we observe exactly k “heads” outcomes from n independent coin flips biased to land on heads with probability p .

The Binomial distribution mass function



Shifting the Binomial distribution

D16-S06(a)

The n Bernoulli trials need not have outcome $+1$ or 0 .

Suppose that the Bernoulli-type random variable \tilde{X} has distribution,

$$\tilde{X} = \begin{cases} b, & \text{with probability } p \\ a, & \text{with probability } 1 - p \end{cases}$$

for any numbers a, b . Then note that if we write

$$\tilde{X} = (b - a)X + a,$$

then $X \sim \text{Bernoulli}(p)$.

Shifting the Binomial distribution

D16-S06(b)

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then $X \sim \text{Bernoulli}(p)$.

Then if we consider a sum of n iid copies of \tilde{X} :

$$\tilde{Y} := \sum_{j=1}^n \tilde{X}_j, \quad \tilde{X}_j \stackrel{\text{iid}}{\sim} X,$$

then we observe that,

$$\tilde{Y} = na + (b - a)Y,$$

where $Y \sim \text{Binomial}(n, p)$ is a (“standard”) Binomial random variable.

Shifting the Binomial distribution

D16-S06(c)

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Hence: iid sums of scaled/shifted Bernoulli random variables are just scaled/shifted Binomial random variables.

Recall that the log-return in our security model satisfies,

$$L = \sum_{j=1}^n L_j,$$

with L_j iid, having distribution,

$$L_j = \begin{cases} \log u, & \text{with probability } p \\ \log d, & \text{with probability } 1 - p \end{cases}$$

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$\log\left(\frac{u}{d}\right) = \log u - \log d$

Hence, we essentially know everything about L :

- L_j is a shifted Bernoulli(p) random variable. Precisely, $L_j = \log d + X_j \log \frac{u}{d}$, where $X_j \sim \text{Bernoulli}(p)$.
- L is a shifted Binomial(n, p) random variable: $L = n \log d + Y \log \frac{u}{d}$, where $Y \sim \text{Binomial}(n, p)$.

(This is essentially why this is called the Binomial pricing model.)

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Some caveats: S_n is not a Binomial random variable. Recall that $S_n = S_0 \exp L$.

The problem is that e^L is not Binomial (although its outcomes have the same probabilities as a Binomial random variable, they are not equispaced outcomes).



Although S_n is not Binomial, it's easy to use knowledge of the Binomial distribution to analyze outcomes.

Example

Consider a 10-period Binomial pricing model, with $(p, u, d) = (0.6, 1.1, 0.9)$.

- Compute the mean and variance of the terminal-time log-return.

$$n=10$$

$$L = \sum_{j=1}^{10} L_j, \quad L_j = \begin{cases} \log 1.1 & \text{w/prob } 0.6 \\ \log 0.9 & \text{w/prob } 0.4 \end{cases}$$

$$L_j = \log 0.9 + X_j \log \left(\frac{1.1}{0.9} \right), \quad X_j \sim \text{Bernoulli}(0.6)$$

$$\mathbb{E} L_j = \log 0.9 + \underbrace{(\mathbb{E} X_j)}_p \log \left(\frac{1.1}{0.9} \right) = \log 0.9 + 0.6 \log \left(\frac{1.1}{0.9} \right)$$

$$\mathbb{E} L = \sum_{j=1}^{10} \mathbb{E} L_j = 10 \cdot \left[\log 0.9 + 0.6 \cdot \log \left(\frac{1.1}{0.9} \right) \right] = 0.1504 \dots$$

$$\text{Var } L = \text{Var} \sum_{j=1}^{10} L_j = \sum_{j=1}^{10} \text{Var } L_j$$

variance of sum of
iid variables is the
sum of variances

$$\begin{aligned} \text{Var } L_j &= \text{Var} [0.9 + X_j \log(\frac{1.1}{0.9})] = \log^2(\frac{1.1}{0.9}) \text{Var } X_j \\ &= \log^2(\frac{1.1}{0.9}) \underbrace{(0.6)(0.4)}_{p(1-p)} \end{aligned}$$

$$\Rightarrow \underline{\text{Var } L} = 10 \cdot \text{Var } L_j = \underline{0.4816}.$$

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Example

Consider a 10-period Binomial pricing model, with $(p, u, d) = (0.6, 1.1, 0.9)$.

- Compute the mean and variance of the terminal-time log-return.
- What is the expected value of S_{10} ?

$S_{10} = S_0 \cdot G$ G is a discrete RV with probabilities given by a Binomial $(10, 0.6)$ RV, with values $u^k d^{10-k}$, for $k=0, \dots, 10$.

$$\mathbb{E} S_{10} = S_0 \mathbb{E} G = S_0 \left[u^0 d^{10} \underbrace{p^0 (1-p)^{10} \binom{10}{0}}_{\substack{\text{Binomial } (10, 0.6) \\ \text{prob.}}} + u^1 d^9 p^1 (1-p)^9 \binom{10}{1} + \dots \right]$$

$$= S_0 \sum_{k=0}^{10} u^k d^{10-k} p^k (1-p)^{10-k} \binom{10}{k} \quad (p, u, d) = (0.6, 1.1, 0.9)$$
$$= 1.219(S_0).$$

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- What is the expected value of S_{10} ?
- What is the probability that $S_{10} \geq S_0$?

$$S_{10} = S_0 \cdot u^x d^{10-x}, \quad x: \# \text{ "heads" outcomes.}$$
$$= S_0 \left(\frac{u}{d}\right)^x d^{10}$$

$$\text{When is } \left(\frac{u}{d}\right)^x d^{10} \geq S_0 \text{ ?}$$

What's the smallest x s.t. $(\frac{1.1}{0.9})^x (0.9)^{10} \geq 1$?

$$x=6$$

$\Rightarrow x=6, 7, 8, 9, 10$ correspond to $S_1 \geq S_0$.

$$\begin{aligned} P(S_1 \geq S_0) &= P(x \geq 6) = \sum_{k=6}^{10} \Pr(x=k) = \sum_{k=6}^{10} \binom{10}{k} (0.6)^k (0.4)^{10-k} \\ &= 0.63 \end{aligned}$$

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- What is the expected value of S_{10} ?
- What is the probability that $S_{10} \geq S_0$?
- What is the probability that $S_5 \geq S_0$?

$$S_5 = S_0 \left(\frac{1.1}{0.9}\right)^x (0.9)^{5-x}, \quad x = 0, 1, 2, \dots, 5$$

$$\left(\frac{1.1}{0.9}\right)^x (0.9)^{5-x} \geq 1 \quad \text{when } x \geq 3$$

$$\Rightarrow \Pr(S_5 \geq S_0) = \sum_{k=3}^5 \binom{5}{k} 0.6^k (0.4)^{5-k} = 0.683$$

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- What is the expected value of S_{10} ?
- What is the probability that $S_{10} \geq S_0$?
- What is the probability that $S_5 \geq S_0$?
- Suppose $S_5 = S_0(1.1)^5$. What is the distribution of S_{10} conditioned on this outcome?

The distribution of S_{10} conditioned on $S_5 = (1.1)^5 S_0$ is the same as the distribution of a 5-period model with initial value $(1.1)^5 S_0$.

$$S_{10} \text{ conditioned on } S_5 = (1.1)^5 S_0 = \begin{cases} (1.1)^5 S_0 (1.1)^5, & \text{w/ prob } p^5 (1-p)^0 \binom{5}{0} \\ \vdots & \vdots \\ (1.1)^5 S_0 (0.9)^5 & \text{w/ prob } p^0 (1-p)^5 \binom{5}{5} \end{cases}$$

(this model is "memoryless")

Some final observations

The process we have defined and investigated has some useful properties:

There are 2^n possible pricing trajectories that the model can take. (2 options at each time.)

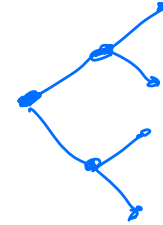
But there are only $j + 1$ possible outcomes at time t_j : the security price S_j is given by,

$$S_j = S_0 \prod_{q=1}^j G_q = S_0 u^{Y_j} d^{j-Y_j},$$

where Y_j is the cumulative number of “heads” realizations in the first j steps:

$$Y_j = \sum_{q=1}^j X_q.$$

Since Y_j has exactly $j + 1$ possible outcomes, then S_j , whose randomness depends explicitly and only on Y_j , has exactly $j + 1$ possible outcomes.



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Related to the above, the tree is *recombining*, meaning that there are trajectories that reach the same state using different pathways. In particular,

- outcome $(G_1, G_2) = (u, d)$
- outcome $(G_1, G_2) = (d, u)$

yield the same security price $S_2 = udS_0$.

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As we saw in the previous example, this is a *Markovian* process: the possibilities at time t_{j+1} can be deduced entirely from the state of things at time t_j .



Petters, Arlie O. and Xiaoying Dong (2016). *An Introduction to Mathematical Finance with Applications: Understanding and Building Financial Intuition*. Springer. ISBN: 978-1-4939-3783-7.